

概率论与随机过程

1. Event and their Probabilities
2. Random Variable
3. Random Vectors
4. Sequences of Random Variables
5. Introduce to Stochastic Processes
6. Stationary Processes
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一、Event and their Probabilities

1. 集合 $A-B$, A , but no B .

$$A-B = A-AB = A\bar{B}, A \cup B = A \cup AB$$

2. ① Classical Probability

有限样本空间, 等概率, $P(A) = \frac{\#A}{\#\Omega}$ $\rightarrow$ 表示度量.
互斥, disjoint / mutually exclusive.

子集有有限可加性 $P(\cup A_i) = \sum P(A_i)$, 这几种都满足

如果事件不同, 则取 n , $\binom{N}{n} n!$ 为所有排列数.

- ② Geometric Probability

$$P(A) = \frac{L(A)}{L(\Omega)} \rightarrow \text{测度}, \text{解投针, 互相等待.}$$

- ③ Frequency

$$F_n(A) = \frac{f_n(A)}{n}, \text{频率就是概率.}$$

3. Probability space.

$$\sigma\text{-algebra} \rightarrow \begin{cases} \Omega \in \mathcal{F} \\ F \in \mathcal{F} \Rightarrow \bar{F} \in \mathcal{F} \\ \cup_n F_n \in \mathcal{F} \end{cases} \text{ 则 } \mathcal{F} \text{ 最小为 } \{\emptyset, \Omega\}$$

probability space $(\Omega, \mathcal{F}, P)$

运算性质:

$$P(A \cup B) = P(A) + P(B) - P(AB)$$

$$P(A \cup B \cup C) = P(A) + P(B) + P(C)$$

$$- [P(AB) + P(AC) + P(BC)] + P(ABC)$$

双数项取减号.

$$P(\cup A_i) = \sum_i P(A_i) - \sum_{i < j} P(A_i A_j) + \sum_{i < j < k} P(A_i A_j A_k)$$

$$- \sum_{i < j < k < l} P(A_i A_j A_k A_l) + \dots + (-1)^{n+1} P(A_1 A_2 \dots A_n)$$

4. Conditional Probabilities.

条件概率: $P(A|B) = \frac{P(A \cap B)}{P(B)}$, 条件概率也是概率.

乘法公式: $P(A_1 A_2 \dots A_n) = P(A_1) P(A_2|A_1) P(A_3|A_1 A_2) \dots P(A_n|A_1 A_2 \dots A_{n-1})$

全概率公式: B_i 互斥, $\cup B_i = \Omega \Rightarrow P(A) = \sum P(B_i) P(A|B_i)$.

贝叶斯公式: $P(B_i|A) = \frac{P(B_i) P(A|B_i)}{\sum P(B_j) P(A|B_j)} \rightarrow$ 条件概率乘法公式

I. Independence of Event.

$$P(A \cap B) = P(A)P(B) \Leftrightarrow P(A|B) = P(A) \Leftrightarrow P(B|A) = P(B)$$

互斥及不独立 (因为已有互斥关系了).

A, B 独立, $\bar{A}, A, \bar{B}, B$, 之间 (A 与 B 类) 两两之间都独立.

① 独立事件独立.

mutually independent: 两两独立并且 $P(A \cap B \cap \dots) = P(A)P(B)P(C) \dots$

pairwise independent: 只有两两独立.

随机变量 (事件) independent (即两两 independent), 那么 $P(A_1 A_2 \dots A_n) = P(A_1) \dots P(A_n)$.

② Bernoulli Trials.

而结果, 互相独立. $P_n(k) = \binom{n}{k} p^k (1-p)^{n-k}$

二. Random Variable.

1. 分布. cumulative distribution function $F(x) = P(X \leq x), \forall x \in \mathbb{R}$. (d.f.).

$$F(x) \left\{ \begin{array}{l} \text{单调递增, } x_1 < x_2, F(x_1) \leq F(x_2). \\ \lim_{x \rightarrow -\infty} F(x) = 0, \lim_{x \rightarrow \infty} F(x) = 1. \\ \text{右连续, } F(x) = F(x^+), \forall x \in \mathbb{R}. \end{array} \right. \left\{ \begin{array}{l} P(X > x) = 1 - F(x). \\ P(x_1 < X \leq x_2) = F(x_2) - F(x_1). \\ P(X < x) = F(x^-) \\ P(X = x) = F(x) - F(x^-) \\ P(X \geq x) = 1 - F(x^-). \end{array} \right.$$

① 离散型 $\rightarrow$ Probability function $p(x)$. (p.f.). 分布律表示定理. d.f. 就是求和.

② 连续型 $\rightarrow$ probability density function (p.d.f.) $P(a \leq X \leq b) = \int_a^b f(x) dx$.

d.f. $F(x) = P(X \leq x) = \int_{-\infty}^x f(x) dx, f(x) = \frac{d}{dx} F(x), \int_{-\infty}^{\infty} f(x) dx = F(\infty) = 1$

< 连续型分布函数求法 > $\hat{=} f(x)$ 因为连续所以 $F(x^+)$ 的 + 或 $P(a < X) P(a \leq X)$ 不必在意是否都一样.

函数情况: 有表格, 直接求, 有公式 $F_Y(y) = P(Y=y) = P(g(X)=y) = P(X=g^{-1}(y)) = P_X(g^{-1}(y))$

$$F_Y(y) = P(Y=y) = P(g(X) \leq y) = P(X \leq g^{-1}(y)) = F_X(g^{-1}(y))$$

$$f_Y(y) = \frac{d}{dy} F_Y(y) = \frac{d}{dy} F_X(g^{-1}(y)) = \frac{d}{d(g^{-1}(y))} F_X(g^{-1}(y)) \frac{d(g^{-1}(y))}{dy} = f_X(g^{-1}(y)) \frac{d(g^{-1}(y))}{dy}$$

2. Mathematical Expectation and Variance.

① 期望 Expectation. (绝对收敛时 $\int |x| f(x) dx < \infty$ 时存在).

$$\mu = E(X) = \begin{cases} \sum_{i=1}^{\infty} x_i P(X=x_i), & \text{discrete.} \rightarrow \text{若 } X \text{ 取值非负, } E(X) = \sum_{k=0}^{\infty} P(X > k) \\ \int_{-\infty}^{+\infty} x f(x) dx = \int_{-\infty}^{+\infty} x dF(x), & \text{continuous.} \end{cases}$$

$$E[af(x)+b] = aE[f(x)] + b.$$

$$E[f(x)] = \begin{cases} \sum_{i=1}^{\infty} f(x_i) P(X=x_i). \\ \int_{-\infty}^{+\infty} f(x) f(x) dx = \int_{-\infty}^{+\infty} f(x) dF(x). \end{cases}$$

如果 $f(\cdot)$ 是线性函数, 那么 $E[f(X)] = f[E(X)]$, 反之, 不真.
② 方差 Variance.

$$\sigma^2 = E[(X - \mu(X))^2] = \text{Var}(X) = \begin{cases} \sum_{i=1}^{\infty} (x_i - \mu)^2 P(X=x_i) \\ \int_{-\infty}^{+\infty} (x - \mu)^2 f(x) dx. \end{cases}$$

$$\text{Var}[af(x)+b] = a^2 \text{Var}[f(x)].$$

$$\text{Var}(X) = E(X^2) - E^2(X), \quad \sigma = \text{SD}(X) = \sqrt{\text{Var}(X)}.$$

方差为 0, X 为 constant.

两点分布, 中点分布, 离散分布存在, 连续分布不一定存在.

③ 收敛性定理的应用.

a) $X \geq 0 \Rightarrow \underline{P(X \geq \varepsilon)} \leq \frac{E(X)}{\varepsilon}$, Markov's inequality. $\rightarrow$ 令 $X = (X - \mu)^2$, $\varepsilon = \varepsilon^2$

b) $\underline{P(|X - \mu| \geq \varepsilon)} \leq \frac{\sigma^2}{\varepsilon^2}$, Chebyshev's inequality.

3. Discrete Random Variables.

① Binomial Distribution. $X \sim B(n, p)$, $x = 0, 1, 2, \dots, n$.

$$P(X=x) = \binom{n}{x} p^x (1-p)^{n-x}, \quad E(X) = np, \quad \text{Var}(X) = np(1-p).$$

当 n 非常大, p 非常小时.

② Geometric Distribution. $X \sim G(p)$, $x = 1, 2, \dots$. $\lim_{n \rightarrow \infty} \binom{n}{x} p^x (1-p)^{n-x} = e^{-\lambda} \frac{\lambda^x}{x!}$, where $\lim_{n \rightarrow \infty} np = \lambda$. 子试验的概率会相对应.

前 $x-1$ 次均失败, 最后一次成功. $P(X=x) = (1-p)^{x-1} p$.

$$E(X) = \frac{1}{p}, \quad \text{Var}(X) = \frac{1-p}{p^2}, \quad F(x) = 1 - (1-p)^x.$$

无记忆性. $P(X > x+y | X > x) = P(X > y)$.

③ Poisson Distribution $X \sim \text{Poi}(\lambda)$, $x = 0, 1, \dots$.

$$P(X=x) = e^{-\lambda} \frac{\lambda^x}{x!}, \quad x = 0, 1, \dots$$

$$E(X) = \text{Var}(X) = \lambda.$$

4. Continuous Random Variables

① Uniform Distribution. $X \sim U(a, b)$

$$f(x) = \begin{cases} \frac{1}{b-a}, & a \leq x \leq b \\ 0, & \text{otherwise} \end{cases} \Rightarrow a=0, b=1 \Rightarrow \text{standard uniform distribution}$$

$$E(X) = \frac{a+b}{2}, \text{Var}(X) = \frac{(b-a)^2}{12}$$

如果 $Y = F_X(X)$, 那么变量 Y 必然在 $[0, 1]$ 上均匀分布.

反过来成立, 一个均匀分布 $Y \sim U(0, 1)$, 通过描述 $X = F_X^{-1}(y)$, 则使得 X 的分布函数为 F_X .

② Exponential Distribution. $X \sim \text{Exp}(\lambda)$

满足 $P(X \leq x) = F_X(x)$.

$$f(x) = \begin{cases} \lambda e^{-\lambda x}, & x > 0 \\ 0, & x \leq 0 \end{cases}$$

$$E(X) = \frac{1}{\lambda}, \text{Var}(X) = \frac{1}{\lambda^2}, \text{记忆技巧: } P(X > t+s | X > s) = P(X > t)$$

③ Normal Distribution. $X \sim N(\mu, \sigma^2)$

$$f(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}, -\infty < x < +\infty$$

高斯极限 $\rightarrow \int_{-\infty}^{+\infty} e^{-x^2} dx = \sqrt{\pi}$, $\mu=0, \sigma=1, \Rightarrow$ standard normal distribution.

对于标准正态分布 $Z \sim N(0, 1)$, $\phi(z) = \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}}, -\infty < z < \infty$ 且 $\Phi(z) = P(Z \leq z)$ 查表

可以使用正态分布来近似 Poisson 和 binomial 分布, 从连续到离散, 一般加一个 0.5 校正, 如:

① 生成 $\begin{cases} B, \text{Poi} \rightarrow \text{count} \\ G, \text{Exp} \rightarrow \text{waiting time} \\ N \rightarrow \text{snms} \end{cases}$

$$P(X=x) = P(0.5 < Y < 20.5)$$

(离散) (连续)

三. Random Vectors. 相同样本空间上的不同随机变量.

1. Random Vectors and Joint Distributions.

Joint d.f. $F(x, y) = P(X \leq x, Y \leq y), -\infty < x, y < +\infty$.

$$P(x_1 < X \leq x_2, y_1 < Y \leq y_2) = F(x_2, y_2) + F(x_1, y_1) - F(x_1, y_2) - F(x_2, y_1) \geq 0$$

marginal distribution: $F_X(x) = \lim_{y \rightarrow +\infty} F(x, y)$, 使用画图法理解.

① Discrete Random Vectors.

$P(x, y) = P(X=x, Y=y)$, $P_X(x) = \sum_y P(x, y)$, $F(x, y) = \sum_{x' \leq x} \sum_{y' \leq y} P(x', y')$, 使用表呈现.

② Continuous Random Vectors.

$$F(x, y) = \int_{-\infty}^y \int_{-\infty}^x f(u, v) du dv, \quad P\{(x, y) \in A\} = \iint_{(x, y) \in A} f(x, y) dx dy$$

$$F_X(x) = \int_{-\infty}^x (\int_{-\infty}^{+\infty} f(x, y) dy) dx \Rightarrow F_X(x) = \lim_{y \rightarrow +\infty} F(x, y), \quad f_X(x) = \int_{-\infty}^{+\infty} f(x, y) dy$$

① 性质

② $F(-\infty, y) = 0, F(x, -\infty) = 0$

$F(x, +\infty) = 1$

③ $\lim_{x \rightarrow x_0^+} F(x, y) = F(x_0, y)$

$\lim_{y \rightarrow y_0^+} F(x, y) = F(x, y_0)$

右连续.

③ $f_{X,Y} = \begin{cases} \frac{1}{ab} & \text{where exists} \\ 0 & \text{otherwise} \end{cases} \Rightarrow$ 但是很少有这样的题目, 对吧, 不是挺麻烦的, 要记住, 不过也得准备推导.

Uniform distribution ($0 < x < a, 0 < y < b$) $P(X,Y \in A) = \frac{L(A)}{ab}$.

2维 normal distribution. $N(\mu_1, \mu_2; \sigma_1^2, \sigma_2^2; \rho)$.

边缘分布是联合分布 $N(\mu_1, \sigma_1^2)$ $N(\mu_2, \sigma_2^2)$.

并不是所有联合事件都拥有 Joint pdf.

2. Independence of Random Variables.

$F(x,y) = F_X(x)F_Y(y) \Rightarrow X, Y$ independent, $X \perp Y$.

$X \perp Y \Leftrightarrow \begin{cases} P(x,y) = P_X(x)P_Y(y), \text{ discrete} \\ f_{X,Y}(x,y) = f_X(x)f_Y(y), \text{ continuous} \end{cases}$

独立充要条件: $\mathcal{I} = \{x,y\}, \text{ scope 为矩形 } \}, f_{X,Y}(x,y) \Rightarrow f_X(x)h(x)f_Y(y)$

X, Y 独立, $f_X(x), h(y)$ 也独立.

多变量.

3. Conditional Distributions. (条件概率也是概率)

①离散型.

$P_{X|Y}(x|y) = P(X=x|Y=y) = \frac{P(X=x, Y=y)}{P(Y=y)} = \frac{P(x,y)}{P_Y(y)}$

$F_{X|Y}(x|y) = \sum_{x \leq x} P_{X|Y}(x|y)$

在 Poisson 分布的基础上对 Poisson 视为样本空间, 再进行二项分布,

得到的分布依旧为 Poisson 分布, 唯一变化的是强度变为 λ , 倒也符合 Poisson 定义.

②连续型.

$f_{X|Y} = \frac{f_{X,Y}}{f_Y(y)}$ if $0 < f_X(x) < +\infty$ for $x \in I$, and 0 otherwise.

只有 $f_X(x) > 0$ 时, 条件概率才有意义, 条件概率是单变量概率, 但是它的 scope I 与 y (可能) 有关, 注意讨论存在和范围.

利用连续条件概率, 可以求 $P(Y \geq y | X = x_0)$ 这样的概率, 通常这法是做不了的, 因为没有考虑极限.

另外, 二维联合分布的边缘概率和条件概率都是分布.

4. One Function of Two Random Variables.

①离散. 有些通过表格呈现, 对角线 $\rightarrow x+y$, $\max(x,y)$, $\min(x,y)$

$P_{X+Y}(n) = \sum_x P_X(x)P_Y(n-x)$, 注意求和的范围, 只有 n 为离散分布 $\frac{n-1}{2}$ 开始, 其余都从 0 开始.

$= P_X * P_Y$

$\sum_{x=1}^{n-1}$

$\sum_{x=0}^n$

ex1. 对于独立 $X \sim B(n_1, p)$, $Y \sim B(n_2, p)$, 由二项式定理 $\sum_{k=0}^n \binom{n_1}{k} \binom{n_2}{n-k} p^k (1-p)^{n-k} = (p + (1-p))^n = 1$

有 $X+Y \sim B(n_1+n_2, p)$. 依然为二项分布.

ex2. 对于独立 Poisson, $X \sim \text{Poi}(\lambda)$, $Y \sim \text{Poi}(\mu)$, $X+Y \sim \text{Poi}(\lambda+\mu)$, 参数相加.

凡这两种 (以及后面的 Gauss), 有确定特征, $B \Rightarrow$ 参数变为 n_1+n_2 , $\text{Poi} \Rightarrow$ 参数变为 $\lambda+\mu$.

② 连续型

a) 正态分布性质: 第一类无偏估计量.

$$F_Z(z) = P(Z \leq z) = P(\varphi(X, Y) \leq z) = P((X, Y) \in A) = \iint_A f_{X,Y}(x, y) dx dy$$

在区域 A 中, 将 z 看作一个 constant.

$$f_{\varphi(X,Y)}(z) = \frac{d}{dz} F_Z(z) = \frac{d}{dz} \iint_A f_{X,Y}(x, y) dx dy$$

正规化方法都要使用换元, 但有一个不成熟, 不技巧.

$$\frac{d}{dz} \int_{-\infty}^{+\infty} dx \int_{-\infty}^{z-x} f_{X,Y}(x, y) dy = \int_{-\infty}^{+\infty} f_{X,Y}(x, z-x) dx$$

只需注意 $f_{X+Y} = f_X * f_Y$, $f_{X+Y}(z) = \int_{-\infty}^{+\infty} f_X(x) f_Y(z-x) dx$ 卷积公式, 其来源推.

对 X, Y 独立, 卷积公式, 卷积 + 移位即可, 卷积不必再麻烦.

独立 $X \sim N(\mu_1, \sigma_1^2)$, $Y \sim N(\mu_2, \sigma_2^2)$ 相加, $X+Y \sim N(\mu_1+\mu_2, \sigma_1^2+\sigma_2^2)$ 依然为正态分布.

b) $\max(X, Y)$

$$\text{利用极值性质. } F_Z(z) = P(\max(X, Y) \leq z) = P(X \leq z, Y \leq z) = F_X(z) F_Y(z).$$

加上独立条件更简单.

c) $\min(X, Y)$

$$\text{极值性质. } F_Z(z) = P(\min(X, Y) \leq z) = 1 - P(\min(X, Y) > z).$$

$$= 1 - P(X > z, Y > z).$$

$$= 1 - (1 - F_X(z) - F_Y(z) + F_Z(z)).$$

$$= F_X(z) + F_Y(z) - F_Z(z).$$

三 Numerical Characteristic of Random Vectors.

① 期望 Expectation.

$$\text{线性性质} \Rightarrow E[h(X, Y)] = \int \int h(x, y) f_{X,Y}(x, y) dx dy.$$

必须有 $h(x, y)$, 将 X, Y 联系起来.

$$E(X+Y) = E(X) + E(Y) \rightarrow E(aX+bY+c) = aE(X) + bE(Y) + c.$$

$$f(X) \leq E(|X|), \text{ 独立} \Rightarrow E(f(X)h(Y)) = E[f(X)]E[h(Y)].$$

$$\text{Cauchy-Schwarz inequality } (E(XY))^2 \leq E(X^2)E(Y^2).$$

$$\text{Var}(X+Y) = \text{Var}(X) + \text{Var}(Y) + 2\text{Cov}(X, Y)$$

$$= \text{Var}(X) + \text{Var}(Y) + 2E[(X-E(X))(Y-E(Y))]$$

$$\text{若 } X \perp Y, \text{Var}(X+Y) = \text{Var}(X) + \text{Var}(Y)$$

$$\text{Var}(aX+bY+c) = a^2\text{Var}(X) + b^2\text{Var}(Y)$$

② Covariance and Correlation.

$$\text{协方差: } \text{Cov}(X, Y) = E[(X-E(X))(Y-E(Y))]$$

$$= E(XY) - E(X)E(Y)$$

$$\text{Cov}(aX+c, bY+d) = ab\text{Cov}(X, Y)$$

$$\text{Cov}(X+Y, Z) = \text{Cov}(X, Z) + \text{Cov}(Y, Z)$$

$$\hookrightarrow \int \int (x-E(X))(y-E(Y)) f_{X,Y}(x,y) dx dy$$

correlation coefficient.

$$\rho(X, Y) = \frac{\text{Cov}(X, Y)}{\sigma_X \sigma_Y}$$

$$\text{消除量纲, 归一化线性关系. } \rho(aX+b, cY+d) = \rho(X, Y)$$

独立/不相关, 反之不成立

独立 \$\Rightarrow\$

uncorrelation

$$\text{Cov}(X, Y) = 0$$

$$E(XY) = E(X)E(Y)$$

$$\text{Var}(X+Y) = \text{Var}(X) + \text{Var}(Y)$$

, 但是, 对于 Normal 独立, 特别!

对于联合分布 \$(X, Y) \sim N(\mu_1, \mu_2; \sigma_1^2, \sigma_2^2; \rho)\$ 来说, \$X \perp Y \Leftrightarrow \rho = 0\$.

IV. Sequences of random Variables.

1. 收敛性特征.

$$\{X_n(\omega), \omega \in \Omega\}_{n=1}^{\infty} \rightarrow \{X_n\}_{n=1}^{\infty}$$

$$\text{分布族: } \{f(x_1, \dots, x_k; n_1, \dots, n_k), \forall k, \forall n_i\}$$

IID ~ independent and identically distribution 独立同分布 \$\Rightarrow\$ 马尔可夫链.

$$\mu_X(n) = E(X_n), \underbrace{E(X_n^2)}_{\text{power function}} = E(X_n^2), \sigma_X^2(n), \underbrace{R_X(n, m)}_{\text{自相关}} = E(X_n X_m)$$

$$C_X(n, m) = \text{Cov}(X_n, X_m) \sim \text{自协方差}$$

$$\left. \begin{aligned} C_{XY}(n, m) &= R_{XY}(n, m) - \mu_X(n)\mu_Y(m) \\ \text{互协方差} & \quad \text{互相关} \quad (\text{cross-}) \end{aligned} \right\}$$

2. 收敛类型

① almost surely, \$X_n \xrightarrow{a.s.} X\$ if \$\{\omega \in \Omega: X_n(\omega) \rightarrow X(\omega) \text{ as } n \rightarrow \infty\}\$ 概率为 1.

收敛性很强, 不可能有某些点不收敛, 但这些都是零测集, 不影响.

② in probability \$X_n \xrightarrow{P} X\$, if \$P(|X_n - X| > \epsilon) \rightarrow 0\$ as \$n \rightarrow \infty\$ for all \$\epsilon > 0\$

$$\text{使用切比雪夫不等式进行证明. } P(|X_n - \mu| > \epsilon) \leq \frac{\text{Var}(X_n)}{\epsilon^2}$$

③ in rth mean, \$X_n \xrightarrow{r} X\$, \$E(|X_n - X|^r) \rightarrow 0\$ as \$n \rightarrow \infty\$, \$r=2\$ 时, 叫均方收敛.

a.s. 收敛 + \$|X_n| < \infty \Rightarrow\$ 收敛, 使用 \$E(|X_n - X|^r)\$ 进行证明.

④ $X_n \rightarrow X$ in distribution, $X_n \xrightarrow{D} X$.

$$P(X_n \leq n) \rightarrow P(X \leq x), n \rightarrow \infty.$$

分布一致, 即依分布收敛.

$$\begin{array}{ccc} (X_n \xrightarrow{a.s.} X) & \searrow & \\ \uparrow \text{r.i.} & & (X_n \xrightarrow{P} X) \longrightarrow (X_n \xrightarrow{P} X) \\ (X_n \xrightarrow{r} X) & \searrow & \end{array}$$

3. The Law of Large Numbers, 大数定律 (LLN)

$$\bar{X}_n = \frac{S_n}{n} = \frac{\sum_{i=1}^n X_i}{n}, S_n = \sum_{i=1}^n X_i$$

$$\left\{ \begin{array}{l} \text{WLLN: } \lim_{n \rightarrow \infty} P(|\bar{S}_n - E(\bar{S}_n)| \geq \varepsilon) = 0 \text{ 或 } \lim_{n \rightarrow \infty} P(|\bar{X}_n - E(\bar{X}_n)| \geq \varepsilon) = 0, \text{ P.} \\ \text{SLLN: } P(\lim_{n \rightarrow \infty} |\bar{X}_n - E(\bar{X}_n)| = 0) = 1, \text{ a.s.} \end{array} \right.$$

Chebyshev's LLN.

均值有界, 方差有界 $\Rightarrow P(|\bar{X}_n - \frac{\sum \mu_i}{n}| \geq \varepsilon) \rightarrow 0, \text{ as } n \rightarrow \infty$ 证明 $P(|\bar{X}_n - \frac{\sum \mu_i}{n}| \geq \varepsilon) \leq ?$ 即代雪夫不等式.

Bernoulli's LLN.

$$P(|\frac{n}{n} - p| < \varepsilon) \rightarrow 1, \text{ as } n \rightarrow \infty, \text{ 频率的稳定值为概率.}$$

4. The Central Limit Theorem 中心极限定理 (CLT)

对于 $S_n = \sum_{i=1}^n X_i$, 我们可以用差公式, 但, 对于 Poisson, Normal, binomial, 最简便的方法是通 (X) (μ, σ^2) (n, p) 确定, 而不是差公式 (虽然也是由差公式证明的).

无论 Levy-Lindenberg CLT, De Moivre-Laplace CLT, Liapunov CLT, 都表明.

$$S_n^* = \frac{S_n - E(S_n)}{\sigma_{S_n}}, S_n = \sum_{i=1}^n X_i$$

$$F_n(x) = P(S_n^* \leq x) \rightarrow \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-\frac{s^2}{2}} ds.$$

那可以用正态分布去近似任何独立重复实验, 只要将 $S \rightarrow S^*$ 标准化.

大数定律让我们知道, n 足够大时, 序列的均值收敛到事件的概率均值. 而中心极限定理告诉我们,

$$P(|\bar{X}_n - E(\bar{X}_n)| \geq \varepsilon) = 1 - P(|\bar{X}_n - E(\bar{X}_n)| < \varepsilon).$$

$$= 1 - P\left(\frac{|\bar{X}_n - E(\bar{X}_n)|}{\sqrt{\text{Var}(\bar{X}_n)}} < \frac{\varepsilon}{\sqrt{\text{Var}(\bar{X}_n)}}\right).$$

$$= 2(1 - \Phi(\frac{\varepsilon}{\sqrt{\text{Var}(\bar{X}_n)}})).$$

序列如例的逼近这个均值.

五. Introduction to Stochastic Processes.

1. 定义与性质.

2. The Distribution Family and the moment Functions

3. The Moments of the Stochastic Processes.

平均可求, 方差都存在.

① Mean, Autocorrelation, Autocovariance.

$$\mu_X(t), \sigma_X^2(t), R_X(t_1, t_2) = E(\cancel{X_{t_1}} X_{t_2}) = E(X_{t_1} X_{t_2}), C_X(t_1, t_2) = \text{Cov}(X_{t_1}, X_{t_2})$$

玩随机过程, 一定要分清那个是随机变量, 哪个不是随机变量.

$E(X_{t_1} X_{t_2})$ 是二阶矩, 交换, 非负三角不等式都满足.

已知 $\left\{ \begin{array}{l} \text{first-order distribution of } X_t \rightarrow \mu_X(t), \sigma_X^2(t) \\ \text{second-order distribution of } X_t \rightarrow \mu_X(t_1, t_2), \sigma_X^2(t_1, t_2), \rho \end{array} \right.$

② Cross-correlation and Cross-covariance.

$$R_{XY}(t_1, t_2) = E(X_{t_1} Y_{t_2})$$

$C_{XY}(t_1, t_2) = 0 \Rightarrow$ uncorrelated, 独立必不相关.

$R_{XY}(t_1, t_2) = 0 \Rightarrow$ orthogonal.

4. Stochastic Analysis.

< continuous > < differentiable >.

连续收敛 - 收敛, 收敛的极限也连续.

$$\lim_{\varepsilon \rightarrow 0} E[(X_{t_0+\varepsilon} - X_{t_0})^2] = 0.$$

$\uparrow$ 等价于 (X_t) 连续.

$R_X(t_1, t_2)$ 连续.

$$R_X(s+t) = E(X'_s X'_t) = \frac{\partial^2 R_X(s, t)}{\partial s \partial t}.$$

积分与微分互逆, 解折互逆无碍.

六. Stationary Processes.

0. 和差化积与积化和差.

一鼓作气

$$\left\{ \begin{array}{l} \sin \alpha + \sin \beta = 2 \sin \frac{\alpha+\beta}{2} \cos \frac{\alpha-\beta}{2} \\ \cos \alpha + \cos \beta = 2 \cos \frac{\alpha+\beta}{2} \cos \frac{\alpha-\beta}{2} \end{array} \right.$$

$$\sin \alpha \cos \beta = \frac{1}{2} (\sin(\alpha+\beta) + \sin(\alpha-\beta))$$

$$\cos \alpha \cos \beta = \frac{1}{2} (\cos(\alpha+\beta) + \cos(\alpha-\beta))$$

$$\sin \alpha \sin \beta = \frac{1}{2} (\cos(\alpha-\beta) - \cos(\alpha+\beta))$$



1. Strict Stationary Process. (SSS).

$$F_X(t_1, \dots, x_n; t_1, \dots, t_n) = F_X(t_1+h, \dots, t_n+h; t_1+h, \dots, t_n+h).$$

IID 序列 $F(x_1, x_2, \dots, x_k) = \prod_{i=1}^k F(x_i)$, 既然 SSS

对 $F(x_n, x_m)$ 时 注意讨论 $m=n$ 的情况

$E(X^2) < \infty \Rightarrow$ second order process.

SSS + second order $\Rightarrow$ WSS. $\left. \begin{array}{l} \mu_X(t) \text{ 为 constant.} \\ R_X(t) \text{ 只与差值有关} \end{array} \right\}$ 一阶和二阶分布独立. 不随时间变化产生变化.

2. Wide Stationary Processes (WSS).

second order + μ_X constant + $R_X(t)$ 只依靠差值 $\Rightarrow$ SSS. (这并不意味着 WSS 的 R_X 推 SSS)

$|R_X(t, s)| \leq \sqrt{R_X(s, s)R_X(t, t)} \leq |R_X(0)|$, 零差为子且最大.

类似性质: $R_X(t, s) = R_X(s, t) = R_X(s+t, t) = R_X(-t, 0)$. 三角不等式.

在针对抽象的随机过程证明 WSS 时, $\left. \begin{array}{l} \text{均值直接为 } \text{constant.} \\ R_X \text{ 的求法不直接求 } R_X(t, t+\tau) \end{array} \right\}$ 随机过程 $\langle \text{随机过程} \rangle$ 则删去一部分独立.

$$E(X_t X_{t+\tau}) = I^2 \sum_{n=0}^{\infty} P(N_t = 2n) - I^2 \sum_{n=0}^{\infty} P(N_t = 2n+1)$$

或 -I

反使用 ~~定理~~ 帮助展开证明.

注意讨论 $t=0$ 和负数. $t' = t + \tau$ 讨论全部定义域 $(-\infty, +\infty)$, 独立, 都要对 t 有关.

$$E(X_t X_{t+\tau}) = E(X_{t'} X_{t'-\tau}) = E(X_{t'} X_{t'+\tau})$$

X_t 周期 T , R_X 同样周期 T ; $R_X(0) = R_X(T) \Rightarrow R_X(\tau) = R_X(\tau+T)$ 则有相差 T 倍一直还有.

$[R_X(t_j - t_k)]$ 非负正定, $\alpha \leq \beta \Rightarrow \alpha \geq 0$.

对于 WSS, $R_X(t)$ 在远处任意, 随机过程均收敛.

3. Joint Stationary Process.

$\{X_t\}$ 和 $\{Y_t\}$ 各自平稳且 $R_{X,Y}(t, t+\tau) = R_{X,Y}(\tau)$

那么 $\{X_t, Y_t\}$ jointly stationary process.

4. Ergodicity of Stationary Process.

每个时间段代表全部时间段 $\rightarrow$ ergodic.

ensemble average, ensemble correlation. 系平均, 系相关, 与统计平均中意义一致.

time average, 时平均 $\langle X_t \rangle = \lim_{T \rightarrow \infty} \frac{1}{T} \int_{-T}^T X_t dt$

time correlation, 时相关 $\langle X_t X_{t+\tau} \rangle = \lim_{T \rightarrow \infty} \frac{1}{T} \int_{-T}^T X_t X_{t+\tau} dt$

一般证明时 $\left. \begin{array}{l} \text{均值为0} \\ \text{系也相等即可} \end{array} \right\} \lim_{T \rightarrow \infty} E \left[\left(\frac{1}{T} \int_{-T}^T X_t dt - \mu_X \right)^2 \right] = E \left[\left(\langle X_t \rangle - \mu_X \right)^2 \right] = 0 \Rightarrow$ ergodic in mean.

$\lim_{T \rightarrow \infty} E \left[\left(\frac{1}{T} \int_{-T}^T X_t X_{t+\tau} dt - R_X(\tau) \right)^2 \right] = E \left[\left(\langle X_t X_{t+\tau} \rangle - R_X(\tau) \right)^2 \right] = 0 \Rightarrow$ ergodic in correlation.

WSS + 时遍历性 (均值 + 相关) = ergodic process.

对于随机相位过程或遍历性, 被多时先打后积比较容易, 当然, 两者没有区别, 都比较好.

均值遍历性: $\lim_{T \rightarrow \infty} \frac{1}{T} \int_0^T (1 - \frac{\tau}{T}) [R_X(\tau) - \mu_X^2] d\tau = 0$.

<随机电报过程> 利用上述 $\frac{1}{T} \int_0^T (1 - \frac{\tau}{T}) [R_X(\tau) - \mu_X^2] d\tau = 0$ 证明均值-ergodic.

5. Power Spectral Density of Stationary Processes.

① 信号能量 = $\int_{-\infty}^{\infty} x^2(t) dt = \int_{-\infty}^{\infty} \frac{1}{2\pi} |F(\omega)|^2 d\omega$

信号功率 = $\frac{1}{2T} \int_{-T}^T x^2(t) dt = \frac{1}{2\pi T} \int_{-\infty}^{\infty} |F(\omega)|^2 d\omega$

(Q) 平均功率 = $\lim_{T \rightarrow \infty} \frac{1}{2T} \int_{-\infty}^{\infty} \frac{1}{2\pi} |F(\omega)|^2 d\omega \Rightarrow S_X(\omega) = \lim_{T \rightarrow \infty} \frac{1}{2\pi} |F_T(\omega)|^2$

对于平稳高斯过程是平均功率. $\hookrightarrow$ 功率谱密度 (PSD).

$Q = \lim_{T \rightarrow \infty} E \left(\frac{1}{2T} \int_{-T}^T x^2(t) dt \right) = \lim_{T \rightarrow \infty} \frac{1}{2T} \int_{-T}^T E(x^2) dt$

$= \frac{1}{2T} \int_{-T}^T R_X(0) dt = R_X(0)$.

$R_X(\tau) \xleftrightarrow{FT} S_X(\omega)$, $Q = \int_{-\infty}^{\infty} S_X(\omega) d\omega = \frac{1}{2\pi} \int_{-\infty}^{\infty} S_X(\omega) e^{j\omega 0} d\omega = R_X(0)$.

($S_X(\omega) \geq 0$, $\int_{-\infty}^{\infty} S_X(\omega) d\omega$ 为有限, 且归一化, 使平均功率为总功率).

② Power Spectral Density and Autocorrelation Function.

<Wiener-Khinchine theorem> $R_X(\tau) \xleftrightarrow{FT} S_X(\omega)$, 将功率谱展开.

$e^{-\beta|\tau|} \leftrightarrow \frac{2\beta}{\beta^2 + \omega^2}$

时域相乘 $\leftrightarrow$ 频域卷积 $\cdot \frac{1}{2\pi}$.

$e^{-\beta|\tau|} \leftrightarrow \frac{2\beta}{\beta^2 + \omega^2}$

$e^{j\omega_0 \tau} \leftrightarrow 2\pi \delta(\omega - \omega_0) \Rightarrow 1 \leftrightarrow 2\pi \delta(\omega)$ 复指数 $\leftrightarrow$ 冲激.

$\delta(t - \tau_0) \leftrightarrow e^{-j\omega \tau_0}$

$g(x)h(x) \leftrightarrow \frac{1}{2\pi} G(\omega) * H(\omega)$.

③ Cross-Power Spectral Density

针对 X, Y 为联合平稳序列.

$S_{XY}(\omega) \leftrightarrow R_{XY}(\tau)$.

依然满足维纳-辛钦定理.

7. Finite Markov Chains.

1. Basic Concepts.

markov property $\rightarrow$ 独立 $\Rightarrow$ 有马尔可夫性即马尔可夫链

时齐 $P(X_{n+1}=j | X_n=i) = P(X_{m+1}=j | X_m=i) \Rightarrow P_{ij}(n, n+1)$
 转移矩阵 $\rightarrow$ stochastic matrix. 时间转移.

2. Higher Order Transition Probabilities and Distributions.

$$P_{i10} = P(X_0=i) \quad P_{ij}(n) = P(X_{n+1}=j | X_n=i).$$

< Chapman - Kolmogorov equation (C-K) >

$$P_{ij}(n+m) = \sum_k P_{ik}(n) P_{kj}(m) \Rightarrow P(n) = P^n, p(n) = p(0) P(n).$$

由已知求出转移事件.

使用乘法公式和线性不断拆分, 再求转移矩阵, 最后出结果.

$p(0)$ 和 P 可以决定这个马尔可夫链所有特性.

又见题目选择最合适的解法, 有些算法即平差法, 有些展开代换法.

3. Invariant Distribution and Ergodic Markov Chain.

$P^n = \begin{pmatrix} \pi \\ \vdots \\ \pi \end{pmatrix}$ 随机过程退化为一个随机变量.

$$\pi = \pi P, \pi \rightarrow \text{invariant distribution.}$$

$\exists \pi, \pi P = \pi$ 为正则矩阵 $\Rightarrow$ ergodic $\Rightarrow$ 极限存在.

$\Downarrow$

$$\text{即 } P_{ij}(n) > 0 \text{ for } \forall i, j \in S.$$

初态为 π 的马尔可夫链是平稳过程.

$$\hookrightarrow P\{X_{n_1}=i_1, \dots, X_{n_k}=i_k\}.$$

$$= P\{X_{n_1+n}=i_1, \dots, X_{n_k+n}=i_k\}.$$

求 π $\left\{ \begin{array}{l} \pi = \pi P \\ \sum \pi_i = 1. \end{array} \right.$

解线性方程组.

极限存在, 正则矩阵. $\Leftrightarrow$ 极限存在, 正则矩阵. 即使我们不知道 π , 如果 $P^n \neq \begin{pmatrix} \pi \\ \vdots \\ \pi \end{pmatrix}$, 也说 P^n 极限不存在.

证明遍历性

收敛存在

收敛不存在 (振荡)

无遍历性

收敛 $\rightarrow$ 遍历性

八. Independent - Increment Processes.

1. 独立增量过程引入.

独立增量过程, 平稳增量.

$X_0=0$ 的独立增量过程完全由 $X_t - X_s$ 决定,

$X_0=0$ 且独立增量.

$$G_X(t, s) = E[(X_t - \mu_X(t))(X_s - \mu_X(s))]$$

$$\stackrel{\text{假设 } 0 < s < t}{=} E[X_t - \mu_X(t) = Y_t] = E(Y_t + Y_s) \quad \text{假设 } 0 < s < t$$

$$\checkmark \quad \text{平稳增量} = E[(Y_t - Y_s)(Y_s - Y_0)] + E(Y_s^2)$$

Y_t 为独立增量

$$= E(Y_t - Y_s)E(Y_s - Y_0) + E(Y_s^2)$$

$$= E(Y_s^2) = E[(X_s - \mu_X(s))^2] = \text{Var}_X(s).$$

即如果 $X_0=0$ 且独立增量 $G_X(t, s) = \text{Var}_X[\min(s, t)]$.

另外 $X_0=0$ 的独立增量过程也是 Markov 过程.

2. Poisson Processes

本课程定义的 Poisson 过程为平稳、独立增量过程, 并且初值为 0, 是特殊的计数过程.

$$P(N_{t+s} = n) = e^{-\lambda t} \frac{(\lambda t)^n}{n!}$$

仍然需要知道的, 还是独立增量过程的特点, $G_X(t, s) = \text{Var}_X[\min(t, s)]$.

当条件概率的条件项为一次独立时, 条件概率就是非条件概率.

Poisson Process 为计数过程, 事件之间的时间间隔为指数分布.

3. Gaussian Processes.

高斯过程不是独立增量过程.

高斯过程在任何时点的线性组合都为 Gauss 过程.

对于平稳高斯过程 $\left\{ \begin{array}{l} \mu_X(t) = \mu_X = \text{constant} \\ G_X(t, t+h) = G_X(h), \text{或理解为自相关和协方差} \end{array} \right.$

都只与差值有关, 因为 $G_X(t, t+h) = R_X(t, t+h) - \mu_X^2(t)$ $\mu_X(t)$ 为常数.

高斯过程 2 阶矩一定存在, 对于 Gauss 过程 $SSS \Leftrightarrow WSS$.

<证明一个过程为 Gauss 过程>

即证明 $\sum_{i=1}^n a_i X_{t_i}$ 为高斯分布, for all $t_1, t_2, \dots, t_n \in T$, for all $a_i \in \mathbb{R}$, all $n > 0$.

<写出 Gauss 分布的概率密度函数 (pdf)>

$$\text{pdf}, f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

$$\text{高斯}, X \sim N(\mu, \Sigma) \quad f_{X(t_1, t_2, \dots, t_n)} = \frac{1}{(2\pi)^{\frac{n}{2}} (\det \Sigma)^{\frac{1}{2}}} e^{-\frac{1}{2} (x-\mu)^T \Sigma^{-1} (x-\mu)}$$

均值向量, 协方差矩阵.

参数向量. 13

4. Brownian Motion and Wiener Processes.

$$\text{Wiener process} \left\{ \begin{array}{l} \text{独立增量} \\ \text{平稳增量 } W_t - W_s = W_{t-s}, \text{ where } 0 \leq s \leq t. \\ W_0 = 0 \\ W_t \sim N(0, \sigma^2 t) \end{array} \right.$$

$$\mu_{W(t)} = 0, \quad C_W(s, t) = R_W(s, t) = \sigma^2 \min(s, t).$$

$$\sum_{k=1}^n a_k W_{t_k} = \sum_{k=1}^n b_k (W_{t_k} - W_{t_{k-1}}) \rightarrow \text{Gauss 过程.}$$

$$\min(s, t) = \frac{s+t - |s-t|}{2} \quad (\text{中点或距离的一半}).$$

Wiener Process 是独立增量过程, 也是 Gauss 过程.